# Lecture 2. Computational methods Markov Chain Monte Carlo, Laplace approximation

18.03.2024-22.03.2024 Instructors: Alina Bazarova, Oleg Filatov. Technical issues: Alexandre Strube

# Why computational methods?

Recall that in our target formula for posterior  $p(\theta)$ 

where  $\theta$  are our parameters the **integral** below can get **really nasty**!

**BUT:** this integral is just a constant! Rewrite  $p(\theta | x)$ although possibly varying over a large range.

### What to do?

$$\vartheta(x) = \frac{p(\theta)p(x \mid \theta)}{\int_{\mathbb{R}} p(\theta)p(X \mid \theta)d\theta}$$

$$=\frac{1}{Z}p(x,\theta)$$
, where Z is just a normalising constant,

Monte Carlo integration.

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from posterior distribution  $p(\theta | X)$ 

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(law of large numbers)

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above problem is to do it through a Markov chain having  $p(\theta | X)$  as its stationary distribution.

This is called Markov chain Monte Carlo.

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### <u>Important</u>: We can construct an MCMC algorithm which will have $p(\theta | X)$ as the stationary





At each time *t* the next state  $\theta_{t+1}$  is chosen by first sampling a candidate *Y* from a *proposal* distribution  $q(. | \theta_t)$  which **depends only on the current state**  $\theta_t$  (or not even that)

At each time t the next state  $\theta_{t+1}$  is chosen by first sampling a candidate Y from a that)

where  $\alpha(\theta, Y) = \min\left(1, \frac{p(Y)p(X \mid Y)q(\theta \mid Y)}{p(\theta)p(X \mid \theta)q(Y \mid \theta)}\right).$ 

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# At each time t the next state $\theta_{t+1}$ is chosen by first sampling a candidate Y from a **proposal**

rejection of all possible candidates Y



Recall  $\alpha(\theta, Y) = \min\left(1, \frac{\pi(Y)q(\theta \mid Y)}{\pi(\theta)q(Y \mid \theta)}\right)$ , and hence

### $\pi(\theta_t)q(\theta_{t+1} \mid \theta_t)\alpha(\theta_t, \theta_{t+1}) = \pi(\theta_{t+1})q(\theta_t \mid \theta_{t+1})\alpha(\theta_{t+1}, \theta_t)$

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**Hint:** one of the  $\alpha$ s in the equality above is equal to 1. Moreover, multiply (1) by  $\pi(\theta_t)$ 

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$$I(\theta_{t+1} = \theta_t)[1 - \int q(Y|\theta_t)\alpha(\theta_t, Y)dY]$$

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# **Metropolis-Hastings sampler II** Recall $\alpha(\theta, Y) = \min\left(1, \frac{\pi(Y)q(\theta \mid Y)}{\pi(\theta)q(Y \mid \theta)}\right)$ , and hence $\pi(\theta_t)q(\theta_{t+1} \mid \theta_t)\alpha(\theta_t, \theta_{t+1})$

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means MCMC has <u>converged</u>. The period before convergence is called <u>burn-in</u>

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The first terms on the left-hand side of (3) and (4) are equal by (2), and the second ones by equality  $\theta_t = \theta_{t+1}$ , therefore

# Hence, once sample from stationary has been obtained, all subsequent samples are going to be from it. This

# **Metropolis-Hastings: how it works in practice**

- **1.** Start at **current position** *X*.
- 2. Propose moving to a **new position** Y using proposal q(Y|X)
- 3. Accept/Reject the new position based on the position's adherence to the data and prior distributions using  $\alpha(X, Y)$ 
  - If you accept: Move to the new position Y. Return to Step 1.
  - Else: Do not move to new position, stay at X. Return to Step 1.
- 4. After a large number of iterations, return all accepted positions.



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3. Proposal has to explore the space efficiently, sometimes it requires to perform experimentation and craftsmanship to construct a good one.

Jupyter notebook 2





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Most typical one: random walk,  $q(Y | \theta) = q(|Y - \theta|)$ .



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Important property: acceptance rate - how frequently the proposal gets accepted. Ideally should be 0.2-0.4



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1. Acceptance too high -> chain mixes slowly. Acceptance too low -> chain stops moving.

2. The larger the variance of the proposal is the lower the acceptance rate is.

3. This can be used during burn-in to reach the desired acceptance rate.



**Example:**  $Y \sim N(\theta_t, s)$ , where N is a normal distribution and s is the custom standard

## Important property: acceptance rate - how frequently the proposal gets accepted.

Instead of updating  $\theta$  en bloc it is often more convenient and computationally efficient to divide  $\theta$  into components  $\{\theta_1 \dots \theta_h\}$  and update them one by one.

This means that instead of  $q(Y|\theta)$  we will have  $q(Y_i|\theta_{-i}, \theta_i)$ , where  $\theta_{-i} = \{\theta_1 \dots \theta_{i-1}, \theta_{i+1} \dots \theta_h\}.$ 

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Acceptance probability will then be  $\alpha(\theta)$ 

Instead of updating  $\theta$  en bloc it is often more convenient and computationally efficient

$$\theta_{-i}, \theta_i, Y_i) = \min\left(1, \frac{\pi(Y_i | \theta_{-i})q(\theta_i | Y_i, \theta_{-i})}{\pi(\theta_i | \theta_{-i})q(Y_i | \theta_i, \theta_{-i})}\right)$$





to divide  $\theta$  into components  $\{\theta_1 \dots \theta_h\}$  and update them one by one.

This means that instead of  $q(Y|\theta)$  we will have  $q(Y_i|\theta_{-i}, \theta_i)$ , where  $\theta_{-i} = \{\theta_1 \dots \theta_{i-1}, \theta_{i+1} \dots \theta_h\}.$ 

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Gibbs sampling uses the property of tractability of all *conditional* posterior distributions to get samples from the unknown *full* posterior distribution of all model variables.

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## Gibbs sampling scheme

Assume we have data  $X \sim p(X | \theta_1, \theta_2)$ 

- 1. Randomly initialize  $\theta_1^{(0)}$  and sample  $\theta_2^{(0)} \sim p(\theta_2 | X, \theta_1^{(0)})$
- 2. For step t = 1, ..., T
  - (a) Sample  $\theta_1^{(t)} \sim p(\theta_1 | X, \theta_2^{t-1})$ (b) Sample  $\theta_2^{(t)} \sim p(\theta_2 | X, \theta_1^t)$

# Laplace approximation

**The idea**: find parameters  $\mu$  and  $\Sigma$  such that  $p(\theta | X) \approx N(\mu, \Sigma)$ 

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Taylor series up to the 2nd term:  $f(\theta) \approx f(\theta)$ 

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Even a crocodile is shorter than this expr

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$$\theta_0 = \theta_{MAP} = \arg\max_{\theta} p(\theta | X)$$

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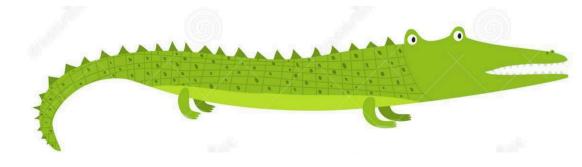
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Hence let us find a good point (MAP):

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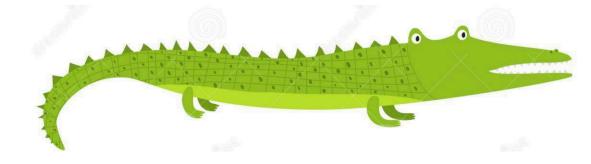
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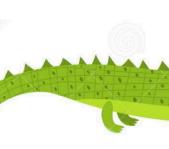
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Hence  $\nabla f(\theta_{MAP}) = 0$  and the second term of the "crocodile" conveniently gets zeroed down:

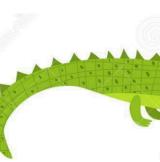


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(Pdf of the normal distribution  $N(\mu, \Sigma)$  is  $p(x, \mu, \Sigma) = (2\pi)$ 

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(Pdf of the normal distribution  $N(\mu, \Sigma)$  is  $p(x, \mu, \Sigma) = (2\pi)$ 

Hence  $\theta | X \sim N(\theta_{MAP}, -(\nabla^2 \ln p(X, \theta_{MAP}))^{-1})$ 

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1. How to find MAP? Iterative procedure, gradient ascent. In **pymc** function **find\_map** which we already used in the first Jupyter notebook.

2. How to find Hessian  $\nabla^2 \ln p(X, \theta)$ ?:

In pymc function find\_hessian

However with the large number of parameters this also becomes too computationally challenging, hence one needs another method

Jupyter notebook 2 Laplace approximation